The Case for Using the General Linear Model as a Unifying Conceptual Framework for Teaching Statistics and Psychometric Theory

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The present paper argues for teaching statistics and psychometric theory using the GLM as a unifying conceptual framework. This helps students understand what analyses have in common, and also provides a firm grounding for understanding that more general cases of the GLM (canonical correlation analysis and SEM) can be interpreted with the same rubric used throughout the GLM. And this approach also helps students better understand analyses that are *not* part of the GLM, such as predictive discriminant analysis (PDA). The approach helps students understand that *all* GLM analyses (a) are correlational, and thus are *all* susceptible to sampling error, (b) can yield r^2 -type effect sizes, and (c) use weights applied to measured variables to estimate the latent variables really of primary interest.

Keywords: General Linear Model, significance testing, effect sizes

In a very influential APA presidential address in the late 1950s, Lee Cronbach advocated greater use of aptitude-treatment interaction designs, and effectively decried the then common misconception that statistics could be conceptualized as fitting within two classes: experimental statistics and correlational statistics. But not much happened with respect to how researchers conceptualized and taught statistics and psychometrics.

Then, in 1968, Cohen published a seminal article that was almost as important as his 1994 article, "The Earth is round (p<.05)." In the 1968 article, Cohen said, although he thought most statisticians would find his argument obvious, most psychologists at the time on the other hand had no idea that regression subsumed ANOVA and other univariate analyses as special cases. Thus, regression is the univariate general linear model (GLM). He argued that the GLM was important conceptually, but also that very important advantages could be realized by using regression to conduct many univariate analyses.

Subsequently, Knapp (1978) showed that canonical correlation analysis (CCA; see Thompson, 1984, 2000) was the multivariate GLM, subsuming in addition to other multivariate methods (e.g., Hotelling T^2 , descriptive discriminant analysis [but not predictive discriminant analysis], and MANOVA and MANCOVA) univariate regression and the other univariate parametric methods. Finally, Bagozzi, Fornell, and Larcker (1981; also see Fan, 1997) showed that structural equation modeling (SEM) was the most general case of the GLM.

A virtual regression-discontinuity study of the influence of Cohen's article shows that the field changed dramatically following Cohen's (1968) publication. Studies by Edgington (1964, 1974) covering several decades showed that prior to the 1968 article around 2/3rds to 3/4ths of published articles used ANOVAs. Similar studies after the 1968 article showed a large drop in the use of ANOVAs (Elmore & Woehlke, 1988; Goodwin & Goodwin, 1985; Kieffer, Reese & Thompson, 2001; Willson, 1980).

The present paper argues for teaching statistics and psychometric theory using the GLM as a unifying conceptual framework. This helps students understand what analyses have in common, and also provides a firm grounding for understanding that more general cases of the GLM (canonical correlation analysis and SEM) can be interpreted with the same rubric used throughout the GLM. And this approach also helps students better understand analyses that are *not* part of the GLM, such as predictive discriminant analysis (PDA). The approach helps students understand that *all* GLM analyses (a) are correlational, and thus are *all* susceptible to sampling error, (b) can yield r^2 -type effect sizes, and (c) use weights applied to measured variables to estimate the latent variables really of primary interest.

More on Defining the GLM

The General Linear Model is the concept that "all analytic methods are correlational ... and yield variance-accounted-for effect sizes analogous to (e.g., R^2 , η^2 , ω^2)" (Thompson, 2000, p. 263). As Graham (2008) explained,

The vast majority of parametric statistical procedures in common use are part of [a single analytic family called] the General Linear Model (GLM), including the *t* test, analysis of variance (ANOVA), multiple regression, descriptive discriminant analysis (DDA), multivariate analysis of variance (MANOVA), canonical correlation analysis (CCA), and structural equation modeling (SEM). Moreover, these procedures are *hierarchical* [italics added], in that some procedures are special cases of others. (p. 485)

Figure 1 presents a conceptual map of the commonly used statistical analyses falling within the General Linear Model. As noted previously, predictive discriminant analysis (PDA), unlike descriptive discriminant analysis (DDA), is *not* part of the GLM (Huberty, 1994). It can also be shown that the mathematics of factor analysis are used to compute the multiplicative weights applied to the measured variables, either explicitly or implicitly, in *all* analyses throughout the GLM.



Figure 1. Conceptual Map of the General Linear Model

Note. Descriptive discriminant analysis (DDA) is part of the general linear model, but predictive discriminant analysis (PDA) is *not* part of the GLM. "SEM" = structural equation modeling; "CCA" = canonical correlation analysis; "*T*-squared" = Hotelling's T^2 , a multivariate extension of the *t*-test to multiple dependent variables.

A very powerful way to prove to students that CCA subsumes other multivariate and univariate parametric methods is to use "proof by SPSS." That is, some students assume that SPSS was written by God, and that therefore anything on an SPSS output must be infallibly true. Here I use the Appendix A heuristic data and the Appendix B SPSS syntax to perform a couple of these proofs via SPSS.

CCA Subsumes Regression as a Special Case

Figure 2 presents a cut-and-paste copy out of an SPSS output for a regression analysis predicting IQ scores with predictors X1, X2, and X3. Figure 3 presents a cut-and-paste copy out of an SPSS output for a CCA with the same measured variables. Note that the regression $R^2 = .04106$ equals the CCA $R_C = .041$.

Figure 2. Cut-and-Paste Copy of SPSS (version 6) Output Regression with IQ as Outcome Variable and Variables X1, X2 and X3 as Predictors

```
* * * * MULTIPLE REGRESSION * * * *
Listwise Deletion of Missing Data
 Equation Number 1 Dependent Variable.. IQ
Block Number 1. Method: Enter
                                                        X1 X2
                                                                                    Х3
 Variable(s) Entered on Step Number 1.. X3
                                                                  X2
                                                        2..
                                                       3.. X1
Multiple R .20263 Analysis of Variance
R Square .04106 DF Sum of Squares Mean Square

        Adjusted R Square
        -.16443
        Regression
        3
        17.42991
        5.80997

        Standard Error
        5.39226
        Residual
        14
        407.07009
        29.07644

                                                      .19982 Signif F = .8948
                                      F =
 ----- Variables in the Equation ------
 Variable
                               В
                                              SE B
                                                              Beta
                                                                                  T Sig T

        X1
        .029676
        .047711
        .178198
        .622
        .5439

        X2
        .016907
        .051648
        .089691
        .327
        .7482

        X3
        .016575
        .046738
        .100762
        .355
        .7281

        (Constant)
        99.489958
        4.497892
        22.119
        .0000
```

Figure 3. Cut-and-Paste Copy of SPSS (version 6) Output CCA with IQ as Predictor Variable and Variables X1, X2 and X3 as Outcomes

```
Eigenvalues and Canonical Correlations

Root No. Eigenvalue Pct. Cum. Pct. Canon Cor. Sq. Cor

1 .043 100.000 100.000 .203 .041

Standardized canonical coefficients for DEPENDENT variables

Function No.

Variable 1

X1 -.879

X2 -.443

X3 -.497
```

However, at first glance the regression beta weights do not appear to match the CCA standardized function coefficients for the parallel analysis. First, the CCA function coefficients each have a different sign than the three regression beta weights! But the scaling direction of equations is purely arbitrary, and any researcher can at will reverse all the signs in an equation within the GLM (see Thompson, 2004, pp. 96-97). This is the equivalent of the arbitrary choice of whether to score a test by counting number of right answers versus number of wrong answers.

Second, the scaling of the regression beta weights and the CCA function coefficients is different. Table 1 illustrates how the two sets of weights can be converted into each other's metrics.

Table 1

Converting Beta Weights into CCA Function Coefficients, and Vice Versa						
	R					
*	0.20263					
*	0.20263					
*	0.20263					
	* * *					

CCA Subsumes Multi-Way ANOVA as a Special Case

Figure 4 presents a cut-and-paste copy out of an SPSS output for an ANOVA summary table for an analysis into IQ as the dependent variable in a two-way factorial ANOVA. Conducting the parallel ANOVA using CCA is a bit

Figure 4. Cut-and-Paste Copy of SPSS (version 6) Output ANOVA with IQ as Outcome Variable and Variables X1, X2 and X3 as Predictor Variables

* * * ANALY	SIS OF	VARIANO	CE * * *			
рλ	IQ EXP_GRP GENDER					
	UNIQUE sums of All effects ea	f squares ntered simulta	aneously			
	Sum o:	£	Mean		Sig	
Source of Variat	ion	Squares	DF	Square	F	of F
Main Effects		412.500	3	137.500	137.500	.000
EXP GRP		300.000	2	150.000	150.000	.000
GENDER		112.500	1	112.500	112.500	.000
2-Way Interactio	ons	.000	2	.000	.000	1.00
EXP GRP GEND	ER	.000	2	.000	.000	1.00
Explained		412.500	5	82.500	82.500	.000
Residual		12.000	12	1.000		
Total		424.500	17	24.971		

tedious, but otherwise is not problematic. First, create orthogonal contrasts using conventional methods explained in various textbooks (e.g., Thompson, 2006). Next, run a CCA model using all the contrast variables. Then run CCA models dropping in turn the contrast variables for the three omnibus effects. Figure 5 presents cut-and-paste copies out of an SPSS output for these analyses.

Figure 5. Cut-and-Paste Copy of SPSS (version 6) Output CCAs with IQ as IQ as Outcome Variable and Variables X1, X2 and X3 as Predictor Variables.

CCA Model #1 wit	h All Five (Orthogonal	Contrasts En	ntered			
EFFECT WITHIN CELLS Regression							
Multivariate	Tests of Si	gnificance	(S = 1, M =	$1 \ 1/2, N =$	5)		
Test Name	Value	Exact F	Hypoth. DF	Error DF	Siq.	of F	
Pillais	.97173	82.50000	5.00	12.00	2	.000	
Hotellings	34.37500	82.50000	5.00	12.00		.000	
Wilks	.02827	82.50000	5.00	12.00		.000	
Rovs	97173						
Note F stat	istics are	exact.					
CCA Model #2 Omi	Ltting Orthod	gonal Contr	asts for the	e Three-Leve	el "A"	Wav	
EFFECT WIT	THIN CELLS R	egression				4	
Multivariate	Tests of Si	gnificance	(S = 1, M =	1/2, N = 6)		
Test Name	Value	Exact F	Hypoth. DF	Error DF	Sia.	of F	
Pillais	.26502	1.68269	3.00	14.00		.216	
Hotellings	.36058	1.68269	3.00	14.00		.216	
Wilks	.73498	1.68269	3.00	14.00		.216	
Rovs	26502	1.00109	0.00	11.00			
Note Estat	istics are a	evact					
Note i Stat	LIDCICS die	chace.					
CCA Model #3 Omi	ltting the O	rthogonal C	Contrast for	the Two-Lev	vel "B	" Way	
CCA Model #3 Omi EFFECT WIT	Itting the Or THIN CELLS Re	rthogonal C egression	Contrast for	the Two-Lev	vel "B	" Way	
CCA Model #3 Omi EFFECT WIT Multivariate	Ltting the Or THIN CELLS Re Tests of Sie	rthogonal (egression gnificance	Contrast for $(S = 1, M =$	the Two-Lev $1 , N = 5$	<u>vel "B</u> 1/2)	" Way	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name	Ltting the Or THIN CELLS Re Tests of Sic Value	rthogonal C egression gnificance Exact F	Contrast for (S = 1, M = Hypoth. DF	the Two-Lev 1 , N = 5 Error DF	<u>vel "B</u> 1/2) Sig.	" Way of F	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais	tting the Or THIN CELLS Re Tests of Sid Value .70671	rthogonal C egression gnificance Exact F 7.83133	Contrast for (S = 1, M = Hypoth. DF 4.00	the Two-Lev 1 , N = 5 Error DF 13.00	<u>vel "B</u> 1/2) Sig.	" Way of F .002	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings	tting the Or THIN CELLS Re Tests of Sid Value .70671 2.40964	rthogonal C egression gnificance Exact F 7.83133 7.83133	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00	the Two-Lev 1 , N = 5 Error DF 13.00 13.00	<u>vel "B</u> 1/2) Sig.	" Way of F .002 .002	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks	CHIN CELLS R Tests of Sid Value .70671 2.40964 .29329	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00	the Two-Lev 1 , N = 5 Error DF 13.00 13.00 13.00	<u>vel "B</u> 1/2) Sig.	" Way of F .002 .002 .002	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys	tting the O: THIN CELLS R Tests of Sid Value .70671 2.40964 .29329 .70671	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00	the Two-Lev 1 , N = 5 Error DF 13.00 13.00 13.00	<u>vel "B</u> 1/2) Sig.	" Way of F .002 .002 .002	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys Note F stat	tting the Or THIN CELLS Re Tests of Sid Value .70671 2.40964 .29329 .70671 tistics are of	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133 exact.	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00	the Two-Lev 1 , N = 5 Error DF 13.00 13.00 13.00	<u>vel "B</u> 1/2) Sig.	" Way of F .002 .002 .002	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys Note F stat	tting the Or THIN CELLS Re Value .70671 2.40964 .29329 .70671 tistics are of tting the Or	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133 exact.	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00	<pre>the Two-Lex 1 , N = 5 Error DF 13.00 13.00 13.00</pre>	vel "B 1/2) Sig.	" Way of F .002 .002 .002 Effects	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys Note F stat CCA Model #4 Omi EFFECT WIT	tting the Or THIN CELLS R Tests of Sid Value .70671 2.40964 .29329 .70671 tistics are of tting the Or THIN CELLS R	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133 exact. rthogonal C egression	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00	<pre>the Two-Lex 1 , N = 5 Error DF 13.00 13.00 13.00</pre>	vel "B 1/2) Sig. action	" Way of F .002 .002 .002 Effects	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys Note F stat CCA Model #4 Omi EFFECT WIT Multivariate	tting the Or THIN CELLS R Tests of Sid Value .70671 2.40964 .29329 .70671 tistics are of thing the Or CHIN CELLS R Tests of Sid	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133 exact. rthogonal C egression gnificance	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00 (S = 1, M =	<pre>the Two-Lex 1 , N = 5 Error DF 13.00 13.00 13.00 r the Intera 1/2, N = 6</pre>	vel "B 1/2) Sig. action	" Way of F .002 .002 .002 Effects	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys Note F stat CCA Model #4 Omi EFFECT WIT Multivariate Test Name	Ltting the Or THIN CELLS R Tests of Sid Value .70671 2.40964 .29329 .70671 Listics are of Ltting the Or THIN CELLS R Tests of Sid Value	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133 exact. rthogonal C egression gnificance Exact F	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00 (S = 1, M = Hypoth. DF	<pre>the Two-Lex 1 , N = 5 Error DF 13.00 13.00 13.00 r the Intera 1/2, N = 6 Error DF</pre>	vel "B 1/2) Sig. action) Sig.	" Way of F .002 .002 .002 Effects	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys Note F stat CCA Model #4 Omi EFFECT WIT Multivariate Test Name Pillais	Ltting the Or THIN CELLS R Tests of Sid Value .70671 2.40964 .29329 .70671 Listics are of Ltting the Or THIN CELLS R Tests of Sid Value .97173	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133 exact. rthogonal C egression gnificance Exact F 160.41667	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00 (S = 1, M = Hypoth. DF 3.00	<pre>the Two-Lex 1 , N = 5 Error DF 13.00 13.00 13.00 r the Intera 1/2, N = 6 Error DF 14.00</pre>	vel "B 1/2) Sig. action) Sig.	" Way of F .002 .002 .002 Effects of F .000	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys Note F stat CCA Model #4 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings	Ltting the Or THIN CELLS Re Tests of Sid Value .70671 2.40964 .29329 .70671 Listics are of Ltting the Or THIN CELLS Re Tests of Sid Value .97173 34.37500	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133 exact. rthogonal C egression gnificance Exact F 160.41667 160.41667	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00 Contrasts for (S = 1, M = Hypoth. DF 3.00 3.00	<pre>the Two-Lex 1 , N = 5 Error DF 13.00 13.00 13.00 r the Inters 1/2, N = 6 Error DF 14.00 14.00</pre>	vel "B 1/2) Sig. action) Sig.	" Way of F .002 .002 .002 Effects of F .000 .000	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys Note F stat CCA Model #4 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks	Ltting the O: THIN CELLS R Tests of Sid Value .70671 2.40964 .29329 .70671 Listics are of Ltting the O: THIN CELLS R Tests of Sid Value .97173 34.37500 .02827	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133 exact. rthogonal C egression gnificance Exact F 160.41667 160.41667	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00 Contrasts for (S = 1, M = Hypoth. DF 3.00 3.00 3.00	<pre>the Two-Lex 1 , N = 5 Error DF 13.00 13.00 13.00 r the Inters 1/2, N = 6 Error DF 14.00 14.00 14.00</pre>	vel "B 1/2) Sig. action) Sig.	" Way of F .002 .002 .002 Effects of F .000 .000	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys Note F stat CCA Model #4 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Boys	Ltting the O: THIN CELLS R Tests of Sid Value .70671 2.40964 .29329 .70671 Listics are of Ltting the O: THIN CELLS R Tests of Sid Value .97173 34.37500 .02827 97173	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133 exact. rthogonal C egression gnificance Exact F 160.41667 160.41667	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00 Contrasts for (S = 1, M = Hypoth. DF 3.00 3.00 3.00	<pre>the Two-Lex 1 , N = 5 Error DF 13.00 13.00 13.00 r the Inters 1/2, N = 6 Error DF 14.00 14.00 14.00</pre>	vel "B 1/2) Sig. action) Sig.	" Way of F .002 .002 .002 Effects of F .000 .000 .000	
CCA Model #3 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys Note F stat CCA Model #4 Omi EFFECT WIT Multivariate Test Name Pillais Hotellings Wilks Roys Note E stat	Ltting the O: THIN CELLS R Tests of Sid Value .70671 2.40964 .29329 .70671 Listics are of Ltting the O: THIN CELLS R Tests of Sid Value .97173 34.37500 .02827 .97173	rthogonal C egression gnificance Exact F 7.83133 7.83133 7.83133 exact. rthogonal C egression gnificance Exact F 160.41667 160.41667	Contrast for (S = 1, M = Hypoth. DF 4.00 4.00 4.00 Contrasts for (S = 1, M = Hypoth. DF 3.00 3.00 3.00	<pre>the Two-Lex 1 , N = 5 Error DF 13.00 13.00 13.00 r the Inters 1/2, N = 6 Error DF 14.00 14.00 14.00</pre>	vel "B 1/2) Sig. action) Sig.	" Way of F .002 .002 .002 Effects of F .000 .000	

Tables 2 and 3 present the calculations to convert CCA lambda values back into conventional ANOVA $F_{CALCULATED}$ values. Obviously, one would not routinely perform ANOVA using CCA, but nevertheless the point that CCA is the multivariate GLM has hopefully been made!

Table 2

Step #1 in Converting CCA Results into Conventional ANOVA $F_{CALCULATED}$ Values: Running Full and Restricted Models for the Various ANOVA Omnibus Effects

						CCA
	Model	Predict	ors			lambda
1	A1	A2	B1	A1B1	A2B1	0.02827
2			B1	A1B1	A2B1	0.73498
3	A1	A2		A1B1	A2B1	0.29329
4	A1	A2	B2			0.02827

Table 3

Step #2: Converting CCA lambda into Ratios into Classical ANOVA $F_{CALCULATED}$ Values

		Full Mode	Lambda	
Effect	Ratio	lambda	w/o effect	Ratio
A Way	1/2	0.02827	0.73498	0.03846
B Way	1/3	0.02827	0.29329	0.09639
AxB Interaction	1/4	0.02827	0.02827	1.00000

Reasons Why I Advocate Teaching Statistics and Psychometrics from a GLM Perspective

Here are some of the reasons why I advocate teaching both statistics and psychometrics from a GLM perspective:

- 1. Teaching statistics and psychometrics from a GLM perspective helps students understand that sampling error effects both $p_{\text{CALCULATED}}$ values and effect sizes no matter what analysis is being done.
- 2. Teaching statistics and psychometrics from a GLM perspective helps students understand that weights are applied to measured variables to estimate latent variables in every analysis. For example, in a balanced ANOVA the eta values are also the beta weights for estimating Y-hat values in ANOVA (see Thompson, 2006).
- 3. Teaching statistics and psychometrics from a GLM perspective helps students understand that it is the design, and *not* the analysis, that

provides the ability to make causal claims (see Thompson, 2006, chapter 12).

- 4. Teaching statistics and psychometrics from a GLM perspective helps students understand that statistics and psychometric models do the same things, albeit it for different purposes: they partition variances (or sum of squares) and estimate ratios of those partitions in forms such as eta squared, R_2 , R_{C2} , and reliability coefficients (see Dawson, 1999).
- 5. Teaching statistics and psychometrics from a GLM perspective helps students understand that all analyses are correlation, which among other things implies that all analyses can be conducted without a researcher's data, as long as one has the covariance matrix and means and *SD*s. These summary statistics are perfectly suitable as inputs into SPSS analyses (see Zientek & Thompson, 2009).

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TEACHING USING THE GLM

Appendix A Hypothetical Heuristic Data

1	1	1	94	790	12	41	14
2	1	1	95	795	51	6	88
3	1	1	96	800	79	99	32
4	1	2	99	780	9	16	79
5	1	2	100	785	60	33	14
6	1	2	101	790	97	30	3
7	2	1	99	785	28	64	90
8	2	1	100	790	46	64	12
9	2	1	101	795	2	23	84
10	2	2	104	775	28	7	31
11	2	2	105	780	84	21	66
12	2	2	106	785	10	73	47
13	3	1	104	800	65	83	39
14	3	1	105	795	75	59	65
15	3	1	106	790	96	69	17
16	3	2	109	770	45	53	91
17	3	2	110	775	47	48	29
18	3	2	111	780	48	42	41

Appendix B SPSS Syntax File for the Illustrative Example

```
SET BLANKS=SYSMIS UNDEFINED=WARN printback=listing .
DATA LIST
 FILE='c:\spsswin\can 14.dta' FIXED RECORDS=1 /1
 id 1-2 exp grp 4 gender 6 ig 8-10 gre 12-14
 x1 16-17 x2 19-20 x3 22-23 .
title 'show CCA is General Linear Model !!! can 14.sps !!!'.
execute .
list variables=all/cases=99999 .
execute .
compute a1 = 0.
if (exp grp eq 1) al = -1.
if (\exp grp eq 2) a1 = 1.
compute a^2 = 0 .
if (a1 ne 0) a2 = -1.
if (\exp grp eq 3) a2 = 2.
if (gender eq 1) b1 = -1.
if (gender eq 2) b1 = 1.
compute al b1 = a1 * b1.
compute a2 b1 = a2 * b1.
print formats al TO a2 b1 (F3) .
list variabes=all/cases=99999 .
execute .
correlations variables = a1 to a2 b1 .
subtitle 'show CCA subsumes regression $$$$$$$$$$$$.
execute .
regression variables=ig x1 TO x3/dependent=ig/
 enter x1 TO x3 .
manova x1 x2 x3 with iq/
 print=signif(multiv eigen dimenr)/
 discrim=stan corr alpha(.999)/.
subtitle 'CCA subsumes factorial ANOVA ======='.
execute .
anova iq by exp grp(1,3) gender(1,2)/statistics=all .
subtitle 'b1 CCA subsumes factorial multi-way ANOVA ||||||'.
execute .
manova a1 a2 b1 a1 b1 a2 b1 with iq/
 print=signif(multiv).
subtitle 'b2 CCA subsumes factorial multi-way ANOVA $$$$$.
execute .
          bl al bl a2 bl with ig/
manova
```

TEACHING USING THE GLM

```
print=signif(multiv).
subtitle 'b3 CCA subsumes factorial multi-way ANOVA @@@@@@'.
execute .
manova al a2 al b1 a2 b1 with iq/
 print=signif(multiv).
subtitle 'b4 CCA subsumes factorial multi-way ANOVA #######'.
execute .
manova al a2 b1
                           with iq/
 print=signif(multiv).
subtitle 'CCA subsumes factorial MANOVA @@@@@@@!.
execute .
manova iq gre by exp_grp(1,3) gender(1,2)/
 print=signif(multiv) .
subtitle 'b1 CCA subsumes factorial multi-way MANOVA ||||||'.
execute .
manova a1 a2 b1 a1 b1 a2 b1 with iq gre/
  print=signif(multiv).
subtitle 'b2 CCA subsumes factorial multi-way MANOVA $$$$$.
execute .
manova
            bl al bl a2 b1 with iq gre/
 print=signif(multiv).
subtitle 'b3 CCA subsumes factorial multi-way MANOVA @@@@@@!.
execute .
manova a1 a2 a1 b1 a2 b1 with iq gre/
 print=signif(multiv).
subtitle 'b4 CCA subsumes factorial multi-way MANOVA #######'.
execute .
manova al a2 b1
                           with iq qre/
 print=signif(multiv).
```